

Corrigendum: On the Precipice of a “Majority-Minority” America: Perceived Status Threat From the Racial Demographic Shift Affects White Americans’ Political Ideology

Craig, M. A., & Richeson, J. A. (2014). On the precipice of a “majority-minority” America: Perceived status threat from the racial demographic shift affects White Americans’ political ideology. *Psychological Science*, 25, 1189–1197. (Original DOI: 10.1177/0956797614527113)

It was recently brought to the authors’ attention that an issue with the base SPSS package leads the “weight by” command, which was used in the analyses reported in this article, to utilize probability weights improperly (i.e., to treat them as frequency weights; for a more comprehensive description of this issue, see UCLA, Institute for Digital Research and Education, 2015). (We are grateful to political scientist L. J. Zigerell for bringing this issue to our attention.) Misspecifying the weighting variable as frequency weights instead of as probability weights resulted in underestimated standard errors and p values in Studies 1 and 2 of this article (Studies 3a and 3b are not affected). The revised and accurate results of these studies are reported here, along with corrected versions of Figure 1, Table 1, and Table 2. Effect sizes remain largely unchanged from those originally reported in the article.

Corrected Results for Study 1

The first two paragraphs of the Results and Discussion section for Study 1 should read as follows:

To analyze the data for this study, we utilized a software package designed for complex survey data (Stata Version 13.1) to ensure that the weighting variable was treated as a probability weight and that subpopulation analyses were conducted accurately.

Logistic regression analyses that regressed party leaning on condition and participants’ demographic characteristics (participants’ gender, education level, and linear and quadratic age terms) revealed a marginal effect of condition, $b = 0.49$, $p = .081$, odds ratio (OR) = 1.63; as predicted, respondents asked about California’s majority-minority racial shift reported leaning somewhat more toward the Republican Party compared with respondents who completed the control survey form. Because the racial shift mentioned was in California, we examined

whether the effect was perhaps stronger in the subsample of respondents living closest to the shift (i.e., in the West Census region; $n = 49$) than in the sample as a whole. As shown in Figure 1, among respondents in the West Census region, respondents’ odds of leaning toward the Republican Party were 3 times as large in the racial-shift condition as in the control condition, $b = 1.17$, $p = .063$, OR = 3.22. Self-reported ideology did not reliably differ by condition, $p > .250$.¹

Corrected Results for Study 2

The first two paragraphs of the Results and Discussion section for Study 2 should be replaced by the following two subsections, and the final paragraph of that section should be preceded by the subhead Discussion.

Weighted results

To test the effect of the racial-shift information on policy issues, we conducted a series of regressions, regressing

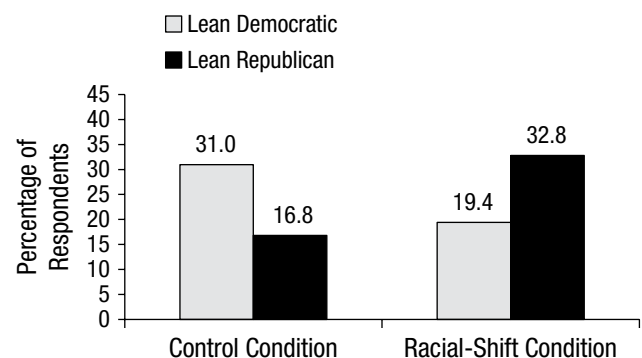


Fig. 1. Results from Study 1: percentage of respondents in the West Census region who leaned Democratic or leaned Republican as a function of condition.

Table 1. Results for Political Attitudes in Study 2: Estimated Marginal Means and Standard Errors by Condition

Measure	Control condition		Racial-shift condition	
	M_{adjusted}	SE	M_{adjusted}	SE
Overall policy attitudes	-0.07	0.04	0.07	0.04
Race-related policy attitudes	-0.07	0.05	0.06	0.05
Race-neutral policy attitudes	-0.08	0.05	0.07	0.05

Note: All items were standardized prior to creating the policy-attitude indices. Means are adjusted for participants' gender, age, and education level. The index of overall policy attitudes included all policy items.

each of the policy indices on experimental condition and participants' demographic characteristics. These analyses were conducted with Stata Version 13.1 and utilized the weighting variable. Results revealed that participants in the racial-shift condition, compared with those in the control condition, expressed more conservative policy positions on the overall policy index, $b = 0.16$, $p = .011$, $r_{\text{partial}} = .14$; the race-related policy index, $b = 0.14$, $p = .052$, $r_{\text{partial}} = .11$; and the race-neutral policy index, $b = 0.20$, $p = .025$, $r_{\text{partial}} = .13$. The only potential mediator that differed significantly by experimental condition was perceived group-status threat, $b = 0.47$, $p = .001$, $r_{\text{partial}} = .17$.

Unweighted results

To be able to test for mediation, we also conducted unweighted analyses. For these analyses, we used SPSS and selected only White participants from the sample; the weighting variable was not utilized.

An analysis of covariance (ANCOVA) predicting overall policy attitudes from experimental condition, controlling for participants' demographics, revealed that participants who read about the racial shift expressed more conservative policy support compared with control participants, $F(1, 464) = 5.84$, $p = .016$, $\eta_p^2 = .01$.⁴ To test

whether this effect was more pronounced for issues directly related to race than for race-neutral issues, we conducted a 2 (political issue: race related, race neutral) \times 2 (condition: control, racial shift) mixed-design ANCOVA, controlling for participants' demographics. Only a main effect of condition emerged, $F(1, 464) = 5.80$, $p = .016$, $\eta_p^2 = .01$ (see Table 1); compared with control participants, participants in the racial-shift condition expressed more support for conservative positions on both race-related issues, $F(1, 464) = 3.82$, $p = .051$, $\eta_p^2 = .01$, and race-neutral issues, $F(1, 464) = 3.89$, $p = .049$, $\eta_p^2 = .01$.⁵

We next examined whether the potential mediators varied by experimental condition, controlling for participants' demographics. The only potential mediator that differed significantly by experimental condition was perceived group-status threat, $F(1, 460) = 15.68$, $p < .001$, $\eta_p^2 = .03$. As predicted, participants in the racial-shift condition perceived more group-status threat ($M_{\text{adjusted}} = 5.31$, $SE = 0.09$) than did participants in the control condition ($M_{\text{adjusted}} = 4.81$, $SE = 0.09$). Furthermore, we analyzed whether perceived group-status threat mediated the effects of experimental condition on policy endorsement. We calculated the indirect effect (Hayes, 2013) of experimental condition on the policy indices through group-status threat, controlling for demographic covariates (with 5,000 bootstrap samples). As shown in Table 2, results were consistent with our hypothesis: Group-status threat mediated the effects of condition on policy support (i.e., bias-corrected confidence intervals do not include 0).

Corrections to Endnotes

The following changes are needed in the endnotes:

- Note 1 should be deleted.
- Note 2 should be renumbered as note 1 and should read as follows: Supplementary analyses of data

Table 2. Mediation Results From Study 2: Path Coefficients and Bias-Corrected Confidence Intervals (CIs)

Coefficient	Overall policy attitudes	Race-related policy attitudes	Race-neutral policy attitudes
a	0.495** (0.125)	0.495** (0.125)	0.495** (0.125)
b	0.094** (0.021)	0.122** (0.024)	0.053† (0.029)
c	0.133* (0.057)	0.125† (0.066)	0.141† (0.077)
c'	0.086 (0.057)	0.065 (0.065)	0.115 (0.078)
95% CI of the indirect effect	[0.020, 0.086]	[0.027, 0.106]	[0.0003, 0.068]

Note: Experimental condition was coded as 1 (racial-shift condition) or 0 (control condition). All models included participants' gender, age, and education level as covariates. Path coefficients are as follows: a denotes the effect of experimental condition on perceived group-status threat; b denotes the effect of perceived group-status threat on the dependent variable; c' denotes the direct effect of experimental condition on the dependent variable; and c denotes the total effect of experimental condition on the dependent variable. Standard errors are in parentheses.

† $p < .10$. * $p < .05$. ** $p < .001$.

from the full White subsample (including Democrats and Republicans) did not reveal an interactive effect of condition and participant's political party or of condition and region (West region, other regions) on expressed political ideology, $ps > .973$.

- Note 3 should be renumbered as note 2 and should read as follows: Supplementary analyses examined the effects of the racial shift on Blacks and Latinos. Among Black participants, those who were asked about the majority-minority shift identified themselves as somewhat more liberal than did those who were in the control condition, $b = -0.415$, $p = .082$; no effect was observed for Latino participants ($p > .355$).
- Note 4 should be renumbered as note 3.
- Note 5 should be renumbered as note 4.
- Note 6 should be renumbered as note 5 and should read as follows: Supplementary analyses examining each policy attitude separately revealed that the experimental manipulation affected attitudes toward affirmative action (weighted $p = .032$; unweighted $p = .015$) and health-care reform (weighted $p = .074$; unweighted $p = .024$), and somewhat affected attitudes toward defense spending (weighted $p = .052$; unweighted $p = .297$). Attitudes toward increasing the number of immigrants (weighted $p = .187$; unweighted $p = .256$) and the required time to be eligible for U.S. citizenship (weighted $p = .501$; unweighted $p = .497$) did not reliably differ by condition.
- Note 7 should be renumbered as note 6 and should read as follows: Supplementary analyses examined the effects of racial-shift salience on Blacks and Latinos. The experimental manipulation had no effects on any dependent measures among Black participants; however, Latinos in the racial-shift condition perceived less system threat (weighted $p = .003$, $r_{\text{partial}} = -.41$; unweighted $p = .064$, $\eta_p^2 = .06$) and expressed more support for defense spending (weighted $p = .010$, $r_{\text{partial}} = .34$; unweighted $p = .125$, $\eta_p^2 = .04$), compared with Latinos in the control condition.
- Note 8 should be renumbered as note 7.
- Note 9 should be renumbered as note 8.

Correction to the Reference List

The article by Graubard and Korn should be deleted.

Corrected Figure and Tables

Figure 1 and Tables 1 and 2 have been updated and are included here where the corresponding corrected results are presented.

Conclusions

In sum, analyses in which the data were appropriately weighted revealed a weaker effect of information about the racial shift on political independents' leanings toward the Republican Party in Study 1. The effect of information about the racial shift on the endorsement of conservative political attitudes in Study 2 remained statistically significant, but with slightly smaller effect-size estimates. Study 3 was not affected by the weighting issue. Thus, its results remain unchanged. Overall, the primary conclusions drawn from this reanalysis are unchanged from those initially drawn: Making projected changes in racial demographics salient for members of the current majority group (White Americans) can lead them to perceive that their racial group's status in society is threatened and, in turn, to express greater political conservatism on both race-related and more race-neutral issues. This work implies that, contrary to suggestions that the changing racial demographics of the United States may be problematic for the Republican Party (e.g., Wickham, 2012), Whites may be increasingly motivated to support conservative candidates and policies in response to these demographic changes, which may lead to an increasingly polarized political landscape.

We regret our error in analyzing the data from Studies 1 and 2, but we hope that this correction can help other researchers who consider analyzing large-scale data sets. We would like to caution researchers seeking to utilize probability weights (the most common type of weights included in large-scale survey data sets) in their analyses that it is important to utilize software that is designed to accommodate complex weighting. The base package of Stata, the survey package for R, and the Complex Samples package (but not the base package) for SPSS can all handle probability weights appropriately.

References

- UCLA, Institute for Digital Research and Education. (2015). *Stata frequently asked questions*. Retrieved from <http://www.ats.ucla.edu/stat/stata/faq/weights.htm>
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